ORIGINAL ARTICLE

Systematic review and meta-analysis of mortality and digoxin use in atrial fibrillation

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Abstract

Background: *There is growing controversy regarding the association between digoxin and mortality in atrial fibrillation (AF). The aim of this analysis was to systematically review digoxin use and risk of mortality in patients with AF.*

Methods: *MEDLINE, EMBASE, GoogleScholar, CINAHL, meeting abstracts, presentations, and Cochrane central databases were searched from inception through December 2014, without language restrictions. For a study to be selected, it had to report the risk of mortality associated with digoxin use in AF patients as an outcome measure. Data were extracted by 2 independent authors. Evidence tables were created.*

Results: *A total of 16 studies (6 post hoc analyses of randomized controlled trials) with 111,978 digoxin users and 389,643 non-digoxin users were included. In a random effects model, patients treated with digoxin had a 27% increased risk of all-cause mortality (pooled HR 1.27; 95% CI 1.19–1.36) and 21% increased risk of cardiovascular mortality (pooled HR 1.21; 95% CI 1.12–1.30) compared with those who did not use digoxin. In a random effects model, the association of digoxin with all-cause mortality was stronger for AF patients without heart failure (pooled HR 1.47; 95% CI 1.25–1.73) than AF patients with heart failure (pooled HR 1.21; 95% CI 1.07–1.36, interaction p = 0.06).*

Conclusions: *Digoxin use in AF is associated with increased risk of all-cause and cardiovascular mortalities. The effect size was larger for AF patients without heart failure than AF patients with heart failure. The study suggests further directed analyses to study the effect that is suggested by this meta-analysis, especially in AF without heart failure.* (Cardiol J 2016; 23, 3: 333–343)

Key words: digoxin, mortality, meta-analysis, systematic review, atrial fibrillation

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Introduction

Digoxin is commonly used for rate control in atrial fibrillation (AF). Despite a relative paucity of data from large randomized controlled trials, several guidelines have recommended its use. The 2014 American Heart Association/American College of Cardiology guidelines suggest the use of digoxin alone in sedentary patients (class I, level of evidence: C) or in combination with either a beta-blocker or non-dihydropyridine calcium channel blocker for non-sedentary individuals and individuals with heart failure (HF) (class IIA, level of evidence: B) [1]. The European Society of Cardiology also recommends its long-term use in both sedentary and active AF patients (class IIA, level of evidence: C) [2]. However, there is growing controversy regarding the association between digoxin and mortality. Earlier investigations of digoxin in HF patients did not demonstrate an increased risk of mortality with digoxin use [3]. However, recent studies suggested an increased risk of mortality with digoxin among AF patients [4, 5]. It will be difficult to design a randomized controlled trial that would evaluate these findings and thus a meta-analysis may provide useful insight and information regarding digoxin use in AF patients. To evaluate the association of digoxin with mortality in AF, we conducted a systematic review and meta-analysis to assess whether digoxin use in AF patients is independently associated with an increased risk of mortality.

Methods

Study design

Protocol. We undertook a systematic review and meta-analysis of randomized controlled trials and observational studies assessing digoxin use in AF patients and its association with mortality.

Data collection. Two reviewers performed an electronic search of MEDLINE, EMBASE, Google-Scholar, Web of Science for published manuscripts until December 15, 2014. We reviewed abstracts from major cardiology meetings held between 2001 and 2014. No limits were used for the search. We also searched the references of the related articles, as well as links to related articles to gather additional articles. We also performed an extensive search of the narrative reviews of the relevant topics. The search terms included variants of "atrial fibrillation", "atrial flutter", "atrial arrhythmia", "supraventricular tachycardia", "digoxin", "digitalis", "cardiac glycoside", "digitoxin", "foxglove", "mortality", "death", "outcomes", "risk factor", "clinical trials", and "prognosis" using text words and Medical Subject Headings (MeSH) terms. We also reviewed editorials and letters related to the topic to identify published and unpublished data. There were no restrictions applied to language, publication date, or publication status. The search was performed without any language restrictions. When an abstract from a meeting and a full article referred to the same trial, only the full article was included in the analysis. When there were multiple reports from the same trial, we used the most complete and/or recently reported data.

Inclusion and exclusion criteria

We included both observational studies and analyses from clinical trials. Randomized controlled trial was defined according to the National Library of Medicine criteria. Atrial fibrillation was defined either by electrocardiogram, self-report or international classification of diseases codes. Studies that did not report mortality were excluded. Data for each trial were abstracted by an investigator (W.T.Q.) and confirmed by a second investigator (M.A.). All discrepancies were identified and resolved by consensus, or as needed, with a third investigator (W.T.O.).

Study quality. We followed the Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA) guidelines for systematic reviews and meta-analyses [6]. We did not use a proper scoring system to grade the study quality as strongly discouraged by the Cochrane Collaboration [7].

Outcomes. The primary endpoint was all-cause mortality. The secondary outcome was cardiovascular mortality. Data on endpoints were abstracted by Waquas Qureshi and Mouaz H. Al-Mallah. All disagreements were resolved by reaching consensus.

Statistical analysis. For all studies, we extracted the study baseline characteristics, event rates, hazard ratios (HR), and 95% confidence intervals (CI) for the primary and secondary outcomes. We extracted HRs for digoxin use in AF patients and all-cause mortality from published manuscripts and abstracts. The effect sizes were obtained from intention-to-treat analyses and fully adjusted models in the cohort studies.

The primary analysis measured the pooled estimate of mortality risk associated with digoxin use in AF patients. The secondary analysis measured the pooled estimate of risk for cardiovascular mortality. A stratified analysis for individuals with

and without HF was also performed as HF patients with AF are more likely to be treated with digoxin due to its inotropic effect. Since we expected significant heterogeneity in the results, we also performed a pre-specified sensitivity analysis without observational studies.

To study heterogeneity, we hypothesized that the effect size may differ according to methodological quality of the studies. Thus, we used a random effects model by DerSimonian and Laird [8]. The random effects model assumes that the studies included in the meta-analysis are a random sample of hypothetical study populations. The random effects model provides a more conservative estimate of the combined data with a wider confidence interval and the summary statistic is less likely to be significant. The heterogeneity was assessed using the Cochrane Q statistic and the percentage of total variability due to true-between study heterogeneity was evaluated by using the I^2 measure. A p-value < 0.10 was considered significant for I^2 measure and interaction tests [9].

We performed meta-regression to examine if the natural log-transformed HR of the effect of digoxin use on mortality was influenced by the prevalence of HF. We used an unrestricted maximal likelihood method for mixed effects regression to evaluate for slope significance.

We assessed publication bias subjectively by visual inspection of Begg's funnel plot [10] and objectively by Egger's regression asymmetry test as funnel plots may be inaccurate in the assessment of very large studies [11, 12]. To address the possibility that "N" number of studies possibly were missing from our analysis and these studies, if included in the analysis, would shift the effect size towards the null, we used Orwin's fail-safe N formula. If the meta-analysis has captured all the relevant studies, then the funnel plot is expected to be symmetrical. However, if there is asymmetry in the plot, it is expected that there are some studies missing from the analysis. This asymmetry is addressed by Duval and Tweedie "trim and fill" method. For example, if there are more studies on the right-hand side and fewer studies on the left-hand side of the funnel plot, this would raise concern that these left-hand studies potentially exist. The Duval and Tweedie trim and fill method trims the right hand asymmetric studies to calculate an unbiased effect by an iterative procedure. It then fills the plot by re-introducing the right hand trimmed studies on the right as well as imputed counterparts to the left of the mean effect [13]. We reported the unbiased effect and the number

of possible missing studies from the analysis. We also performed an additional cumulative analysis to evaluate if there was a temporal effect of the studies. Additionally, we evaluated removal of individual studies on the pooled HR. All analyses were performed using RevMan v. 5.0 and Comprehensive Meta-Analysis v. 2.2.4.

Results

We identified 1,543 studies during the initial literature search. After removing duplicates, 1,242 studies remained. After reading the abstracts and full text of the selected studies, 1,052 studies were discarded due to inability to meet inclusion criteria. Full review of 101 manuscripts and abstracts was performed and 24 studies were selected, 6 did not report the desired HRs and effect sizes to estimate our primary outcome. One of the studies was a patient-level meta-analysis of 4 trials [14]. Finally, 16 studies were selected with study data of 19 studies in this meta-analysis (Suppl. Fig. 1).

As mentioned above, there were 6 studies out of 16 that reported data from 9 clinical trials while other studies were observational in nature. There was 1 conference presentation and 1 conference abstract, while the remaining studies were either published or in online print. We used both of them separately. Study characteristics are presented in Table 1. Studies mainly included males from white populations. Hypertension was the most frequent risk factor (39–100%) and only 1 study did not include patients with HF.

There were a total of 501,681 participants (mean age 73.8 years, males 62.3%), of whom 43,370 were enrolled in clinical trials and 458,311 included in observational studies. There were 414,116 patients with AF. There were 111,978 digoxin users and 389,643 participants who did not use digoxin. A breakdown for each study is presented in Supplementary Table 1.

Using a random effects model, the risk of mortality was 27% higher in individuals using digoxin compared with persons who did not use digoxin (pooled HR: 1.27; 95% CI 1.19–1.36, p < 0.001) (Fig. 1). When only subgroups of randomized clinical trials were analyzed, the risk for mortality for digoxin users increased to 46% (pooled HR: 1.46; 95% CI 1.09–1.94, p = 0.01) (Fig. 2). There was a high degree of heterogeneity in pooled studies (χ^2: 124, df = 14, p < 0.001; I^2 = 89%). When only studies with older individuals (n = 8; age > 70) who had a higher prevalence of HF (> 30%) were retained in the analysis, the heterogeneity became

Table 1. Characteristics of the included studies.

Study	Year	Mean age [years]	Males [%]	Hypertension [%]	Diabetes [%]	Beta-blocker use [%]	ACE inhibitor/ARB use [%]	HF [%]	IHD [%]
LIFE [28]	1997	67	42	100	13	49	50	1.4	15.3
AFFIRM [15]	2001	70	61	71	20	58	64	8.2	38.1
SPORTIF [29]	2002	71	69	77	24	4	NA	36.6	44.7
RIKS-HIA [4]	2003	68	58	39	25	50	35	37.8	55.9
SCAF [5]	2007	77	55	48	18	50	38	44.7	19.1
Fauchier [23]	2007	74	61	45	17	35	73	43	23
TREAT-AF [30]	2008	72	98	63	29	NA	NA	15.6	4.5
ATRIA-CVRN [24]	2009	71	56	78	24	58	44	0	3.7
RACE II [31]	2009	68	66	61	11	70	51	34.5	17.9
ROCKET-AF [32]	2010	73	60	90	40	65	67	54.4	23.3
NHIRD [33]	2010	67	54	66	30	NA	NA	10.9	10.1
PALLAS [34]	2010	75	65	85	NA	41	77	54	41
ADONIS [14]	2010	64	69	54	NA	65	41	12.2	24.2
ANDROMEDA [14]	2010	72	79	39	NA	71	83	100	63.3
ATHENA [14]	2010	72	53	86	NA	74	70	21.2	30
ORBIT-AF [35]	2011	68	57	NA	NA	NA	NA	13	16
Quebec study [27]	2012	80	42	62	32	51	61	32.3	NA
AUPD study [36]	2012	73	53	20	8	18	NA	10	9.4
Pastori study [37]	2013	73	57	89	20	41	70	16	23

ACE/ARB — angiotensin converting enzyme/angiotensin receptor blocker; HF — heart failure; IHD — ischemic heart disease; NA — not available; LIFE — The Losartan Intervention For Endpoint reduction in hypertension; AFFIRM — the AF Follow-Up Investigation of Rhythm Management; SPORTIF — the Stroke Prevention using an ORal Thrombin Inhibitor in atrial Fibrillation; RIKS-HIA — the Registry of Information and Knowledge about Swedish Heart Intensive care Admissions; SCAF — the Stockholm Cohort-Study of Atrial Fibrillation; TREAT-AF — The Retrospective Evaluation and Assessment of Therapies in AF); ATRIA-CVRN — the AnTicoagulation and Risk factors In Atrial fibrillation-Cardiovascular Research Network; RACE II — Rate Control Efficacy in Permanent AF; ROCKET-AF — the rivaroxaban once daily, oral, direct factor Xa inhibition compared with vitamin K antagonism for prevention of stroke and embolism trial in atrial fibrillation; NHIRD — the National Health Insurance Research Database; PALLAS — The Permanent Atrial FibriLLation Outcome Study using dronedarone on top of standard therapy; ADONIS — The American–Australian–African trial with DronedarONe In atrial fibrillation or flutter Patients for the maintenance of Sinus rhythm; ANDROMEDA — the ANtiarrhythmic trial with DROnedarone in Moderate-to-severe CHF Evaluating morbidity DecreAse; ATHENA — A placebo-controlled, double-blind, parallel arm Trial to assess the efficacy of dronedarone 400 mg b.i.d. for the prevention of cardiovascular Hospitalization or death from any cause in patiENts with Atrial fibrillation/atrial flutter; ORBIT-AF — the Outcomes Registry for Better Informed Treatment of Atrial Fibrillation; AUPD — Aarhus University Prescription Database
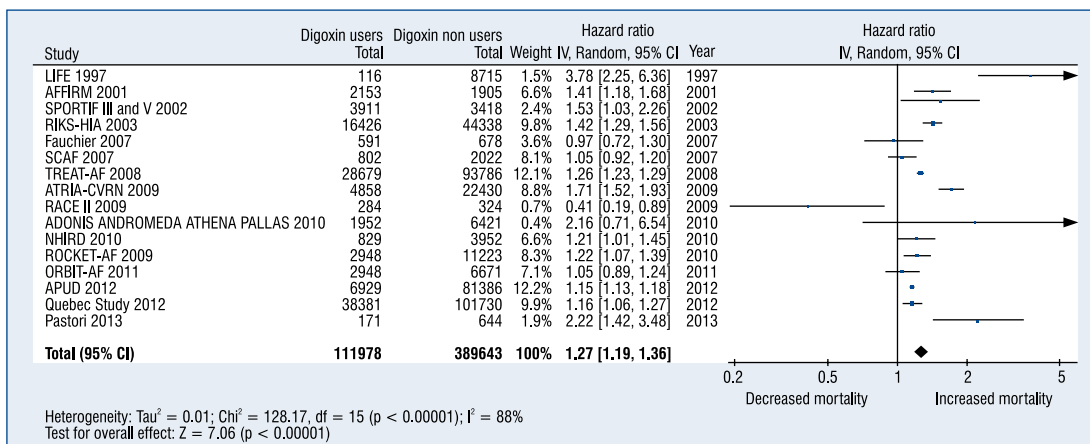

Figure 1. Forest plot of the included studies. The computed weighted hazard ratios are depicted as boxes proportional to the size of the study. The bars depict the 95% confidence intervals (CI). Studies favoring use of digoxin are on the right hand side of the center line of no effect, while the studies against the use of digoxin are on the left hand side of the center line of no effect.

Study of subgroup	Digoxin users Total	Non digoxin users Total	Weight	Hazard ratio IV, Random, 95% CI
SPORTIF III and V 2002	3911	3418	22.6%	1.53 [1.22, 1.92]
ROCKET-AF 2009	2948	11223	24.8%	1.22 [1.06, 1.40]
RACE II 2009	284	324	9.1%	0.41 [0.19, 0.89]
LIFE 1997	116	8715	14.1%	3.78 [2.25, 6.36]
AFFIRM 2001	2153	1905	24.1%	1.41 [1.19, 1.67]
ADONIS ANDROMEDA ATHENA PALLAS 2010	1952	6421	5.3%	2.16 [0.71, 6.54]
Total (95% CI)	**11364**	**32006**	**100.0%**	**1.46 [1.09, 1.94]**

Heterogeneity: Tau^2 = 0.08; Chi^2 = 28.43, df = 5 (p < 0.0001); I^2 = 82%
Test for overall effect: Z = 2.57 (p = 0.01)

Figure 2. Forest plot of subgroup analyses of randomized controlled clinical trials (sensitivity analysis). The computed weighted hazard ratios are depicted as boxes proportional to the size of the study. The bars depict the 95% confidence intervals (CI). Studies favoring use of digoxin are on the right hand side of the center line of no effect, while the studies against the use of digoxin are on the left hand side of the center line of no effect.
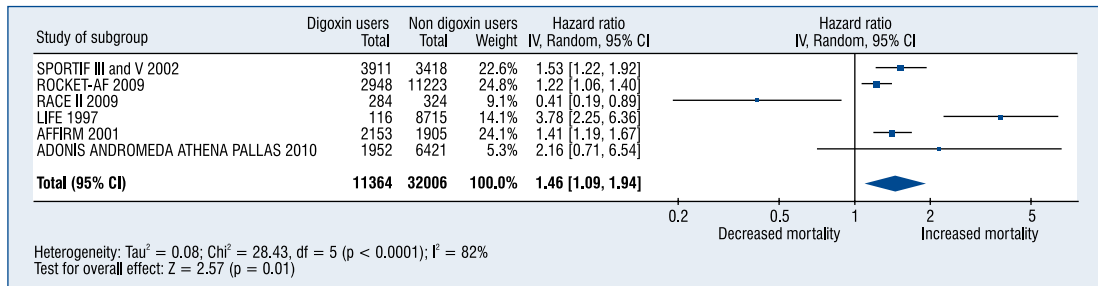
insignificant (χ^2: 7.10, df = 6, p = 0.31; I^2 = 16%). The AF Follow-Up Investigation of Rhythm Management (AFFIRM) trial had 2 published analyses. For primary analysis, we used the study with the larger subsample [15], however in a sensitivity analysis with the smaller sub-study [16], there was no significant difference in the pooled HRs (1.27; 95% CI 1.19–1.36 vs. 1.20; 95% CI 1.18–1.22, p for interaction = 0.11). Other subgroup analyses are shown in Supplementary Table 2. The risk of cardiovascular mortality associated with digoxin use was available in 9 studies. In a random effects model, digoxin use was associated with a 21% increased risk of mortality (pooled HR: 1.21; 95% CI 1.12–1.30, p < 0.001) with significant inter-study heterogeneity (I^2 = 74%, p < 0.001) (Suppl. Fig. 2). Removal of studies with < 30% of HF showed no heterogeneity (p > 0.1).

In a stratified analysis, the risk of mortality for individuals using digoxin in AF patients with HF was 21% higher (pooled HR: 1.21; 95% CI 1.07–1.36, p = 0.002; Fig. 3A) compared with participants who did not use digoxin. The risk of mortality for individuals without HF was 47% higher for digoxin users (pooled HR: 1.47; 95% CI 1.25–1.73, p < 0.001; Fig. 3B) than non-digoxin users.

Due to heterogeneity, a funnel plot was created to assess publication bias (Suppl. Fig. 3). The funnel plot showed minimal asymmetry that was quantified statistically by Egger's regression intercept (intercept: 1.16; 95% CI –0.76–3.08, df = 14, 1-tailed p-value: 0.11). Classical fail safe N was 1507, suggesting that there would need to be 1507 null studies with mean HR of 1.00 added to the analysis before the pooled effect would become non-significant. A more conservative method, Orwin's fail-safe N method was used to estimate the number of studies that would need to be added to the meta-analysis to make the pooled effect non-significant. For this reason, we assumed a mean HR of less than 0.90 for the potentially missing studies. Using this criterion, at least 29 studies would be needed to make the association between digoxin use and mortality non-significant. Using a trim and fill method, we needed to trim a study on the left of the mean which changed the pooled estimate from 1.28 (95% CI 1.20–1.37) to 1.26 (95% CI 1.18–1.35). There was no asymmetry on the right of the mean.

Additional analyses showed that there was a temporal effect present with more recent studies having lower effect sizes than older studies, however the direction and point estimate for the pooled HRs remained significant (Suppl. Fig. 4). When we removed one study from the analysis at each step of the analysis, the pooled estimate remained significant (Suppl. Fig. 5). In meta-regression analysis, proportion of HF in studies was inversely associated with risk of mortality associated with digoxin (slope: –0.008; 95% CI from –0.01 to –0.002, p = 0.02, Suppl. Fig. 6) suggesting HF modifies the risk of mortality and digoxin use.

Discussion

This comprehensive meta-analysis included 501,681 individuals and assessed the effect of digoxin use in AF patients in 9 randomized clinical trials and 10 observational cohorts. There were three main findings: 1) digoxin users with AF had a 27% increased risk of all-cause mortality, 2) digoxin users with AF had a 21% increased risk of cardiovascular mortality, and 3) digoxin users with AF who did not have HF had a relatively stronger association with risk of all-cause mortality (47% vs. 21%, p for interaction = 0.058) than digoxin users with HF.
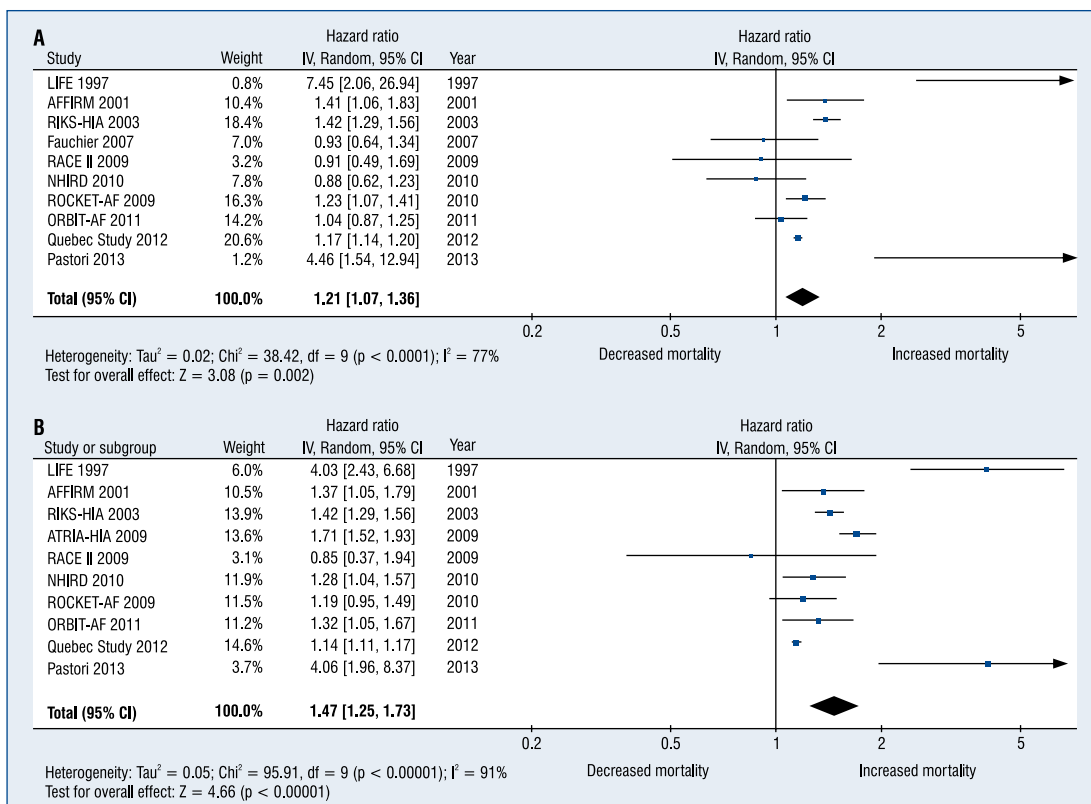
Figure 3. Forest plot of studies including atrial fibrillation patients with heart failure showing association of digoxin use with all-cause mortality. The computed weighted hazard ratios are depicted as boxes proportional to the size of the study. The bars depict the 95% confidence intervals (CI). Studies favoring use of digoxin are on the right hand side of the center line of no effect, while the studies against the use of digoxin are on the left hand side of the center line of no effect; **A.** Heterogeneity becomes insignificant after removing, LIFE, RIKS-HIA, RACE-II and Pastori study, $I^2 = 40\%$, p = 0.14; **B.** Heterogeneity becomes insignificant after removing, LIFE, ATRIA-CVRN, RIKS-HIA and Pastori study, $I^2 = 1\%$, p = 0.41.

Several reports have examined the negative consequences of digoxin use resulting in a significant drop in digoxin use in acute myocardial infarction during the last two decades [17, 18]. Data supporting its use in chronic systolic HF have mainly come from two withdrawal studies [19, 20], and the Digitalis Investigation Group trial that showed an absolute increase of 1.9% in cardiac death with digitalis (not including HF-related death) [21]. However, due to concomitant absolute reduction in HF-related mortality by 1.6%, the overall primary endpoint was not significant.

The aforementioned studies found an increased risk of all-cause mortality in conditions that are not limited to AF. Our results support these prior findings and extend prior work to show an increased risk of all-cause and cardiovascular mortalities exist for AF patients who use digoxin. Although the risks for the adverse outcomes examined were attenuated when we examined patients with HF, there was still a suggestion of 21% increased risk of all-cause mortality in these patients.

We also observed a downward temporal trend in reduced risk of mortality associated with digoxin use over the last 15 years. The reason for this difference is speculative but possibly related to the improved management of cardiovascular disease over the last decade [22]. Several of these studies have included individuals over periods of time, and have found a decreasing trend towards reduced prescription of digoxin by the clinicians [23].

The exact mechanism underlying increased mortality risk observed with digoxin is unknown. Most of the reviewed studies did not evaluate the serum digoxin levels with mortality. Serum digoxin levels were significantly higher in those who died (1.151 vs. 0.935 ng/mL, p < 0.001) in one of the studies [24]. Renal function regulates serum digoxin levels and a subgroup analysis of AF-FIRM demonstrated increased mortality in digoxin

users (HR: 1.28; 95% CI 1.25–1.31; p < 0.001). Thus it can be speculated that higher serum levels of digoxin are related with the pathogenesis. Sudden cardiac death was also observed to be high in ROCKET-AF study and LIFE study in digoxin users. Sudden cardiac death is frequently due to fatal arrhythmia [25]. Digoxin was shown to increase arrhythmia-related deaths by 61% in *post hoc* analysis of AFFIRM trial [15]. Furthermore, we identified association of digoxin with cardiovascular mortality as well suggesting this observed mortality has cardiac basis rather than non-cardiovascular cause. This has been also previously demonstrated in a study without patients with AF [15]. Digoxin was shown to increase arrhythmia-related deaths by 61% in *post hoc* analysis of AFFIRM trial [15].

This study has several implications regarding management of AF in patients with and without HF. The burden of AF is increasing at dramatic pace with doubling AF by 2050. The annual mortality risk in AF is 3.84%, of which 37.4% occur due to cardiovascular conditions [26]. This analysis shows that at least $1/4^{th}$ of this risk is probably related to digoxin. A possible reduction in prescription of digoxin to these individuals may potentially improve overall survival of these patients. Even though the results have been consistent, this study still is limited by its exploratory design and thus calls for a randomized trial of digoxin in AF patients, with and without HF, to demonstrate its safety. However, it may be difficult to perform a randomized study at this point; a withdrawal study might be more suitable to confirm our findings. Additionally, this study suggests that there is a need to revisit the management guidelines of AF. There have been several changes in the management of HF and AF in the last decade. Management of AF in the light of recent evidence-based therapies may eventually lead to discarding the use of digoxin in these particular patients.

Limitations of the study

The current analysis should be interpreted in the context of several limitations. We did not have access to the source data for any of the studies used and the analysis was based on effect sizes and confidence intervals obtained from published studies. There was a significant heterogeneity in the population which remained unexplained and possibly is due to differences in the study samples examined. Also, the included studies from clinical trials were performed as *post hoc* analyses and potentially introduced bias into the current study. Bias by indication also is a significant limitation

that should be considered while interpreting these results. Additionally, we were not able to assess the type of AF (e.g., paroxysmal, persistent, permanent) and the results may vary when accounting for this aspect of the arrhythmia. The exact doses of digoxin and other AF medications were unknown and possibly influenced the relationship between digoxin and the outcomes examined (e.g., drug-drug interactions) [27].

Conclusions

The results from this pooled analysis suggest that digoxin use in AF is associated with an increased risk of all-cause and cardiovascular mortalities. Special consideration should be used by clinicians who use digoxin for rate control in patients with AF. Further study is needed to elucidate the underlying mechanism related to the increased risk of adverse events associated with digoxin.

Conflict of interest: None declared

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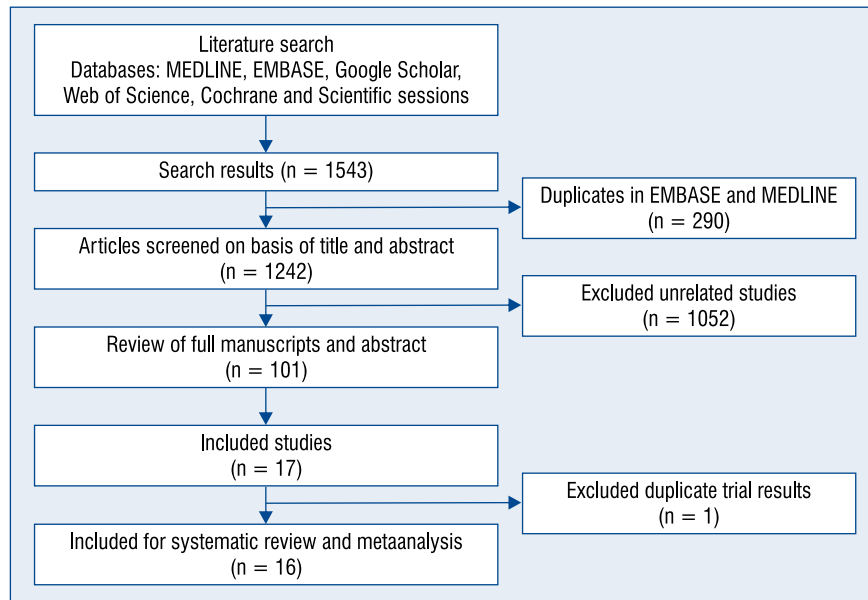
Supplementary Figure 1. Flowchart of the included studies.

**Supplementary Table 1.** Number of participants in the clinical studies and follow-up duration.

Study	Year	Digoxin users	Non digoxin users	Atrial fibrillation	Deaths	Total	Follow-up
LIFE [28]	1997	116	8,715	701	151	8,831	4.7 years
AFFIRM [15]	2001	2,153	1,905	4,058	666	4,058	3.5 years
SPORTIF [29]	2002	3,911	3,418	7,329	396	7,329	1.5 years
RIKS-HIA [4]	2003	16,426	44,338	60,764	7,520	60,764	1 years
SCAF [5]	2007	802	2,022	2,824	1,038	2,824	4.6 years
Fauchier [23]	2007	591	678	1,269	247	1,269	881 days
TREAT-AF [30]	2008	28,679	93,786	122,465	28,723	122,465	0.4 years
ATRIA-CVRN [24]	2009	4,858	22,430	27,288	1,140	27,288	1.17 years
RACE II [31]	2009	284	324	608	76	608	3 years
ROCKET-AF [32]	2010	2,948	11,223	14,171	1,214	14,171	3 years
NHIRD [33]	2010	829	3,952	3,952	170	4,781	2.3 years
PALLAS [34]	2010	1,070	2,166	3,236	31	3,236	117 days
ADONIS [14]	2010	118	511	629	13	629	370 days
ANDROMEDA [14]	2010	135	105	240	17	240	228 days
ATHENA [14]	2010	629	3,639	4,268	153	4,268	651 days
ORBIT-AF [35]	2011	2,948	6,671	9,619	178	9,619	3 years
Quebec study [27]	2012	38,381	101,730	140,111	79,312	140,111	3 years
AUPD study [36]	2012	6,929	81,386	8,880	12,826	88,315	365 days
Pastori study [37]	2013	171	644	815	85	815	33.2 months

LIFE — The Losartan Intervention For Endpoint reduction in hypertension; AFFIRM — the AF Follow-Up Investigation of Rhythm Management; SPORTIF — the Stroke Prevention using an ORal Thrombin Inhibitor in atrial Fibrillation; RIKS-HIA — the Registry of Information and Knowledge about Swedish Heart Intensive care Admissions; SCAF — the Stockholm Cohort-Study of Atrial Fibrillation; TREAT-AF — The Retrospective Evaluation and Assessment of Therapies in AF); ATRIA-CVRN — the AnTicoagulation and Risk factors In Atrial fibrillation-Cardiovascular Research Network; RACE II — Rate Control Efficacy in Permanent AF; ROCKET-AF — the rivaroxaban once daily, oral, direct factor Xa inhibition compared with vitamin K antagonism for prevention of stroke and embolism trial in atrial fibrillation; NHIRD — the National Health Insurance Research Database; PALLAS — The Permanent Atrial FibriLLation Outcome Study using dronedarone on top of standard therapy; ADONIS — The American–Australian–African trial with DronedarONe In atrial fibrillation or flutter Patients for the maintenance of Sinus rhythm; ANDROMEDA — the ANtiarrhythmic trial with DROnedarone in Moderate-to-severe CHF Evaluating morbidity DecreAse; ATHENA — A placebo-controlled, double-blind, parallel arm Trial to assess the efficacy of dronedarone 400 mg b.i.d. for the prevention of cardiovascular Hospitalization or death from any cause in patiENts with Atrial fibrillation/atrial flutter; ORBIT-AF — the Outcomes Registry for Better Informed Treatment of Atrial Fibrillation; AUPD — Aarhus University Prescription Database

Supplementary Table 2. Subgroup analyses by various characteristics of studies.

	Risk of mortality Hazard ratio (95% CI)		P*	Risk of cardiovascular mortality Hazard ratio (95% CI)		P*
Age	> 70: 1.20 (1.18–1.21)	≤ 70: 1.33 (1.24–1.42)	0.003	> 70: 1.17 (1.05–1.30)	≤ 70: 1.26 (1.13–1.41)	0.34
Follow up	> 2 years: 1.22 (1.17–1.28)	≤ 2 years: 1.20 (1.18–1.22)	0.67	> 2 years: 1.21 (1.12–1.31)	≤ 2 years: 1.20 (0.88–1.63)	0.05
BB use	> 50%: 1.21 (1.13–1.29)	≤ 50%: 1.20 (1.18–1.22)	0.22	> 50%: 1.20 (1.10–1.30)	≤ 50%: 1.33 (1.06–1.68)	0.44
Year 2010	During and after: 1.16 (1.14–1.18)	Before: 1.28 (1.25–1.30)	< 0.001	During and after: 1.23(1.11–1.37)	Before: 1.18 (1.06–1.32)	0.59
IHD	> 20%: 1.36 (1.27–1.45)	≤ 20: 1.19 (1.18–1.21)	< 0.001	> 20%: 1.29 (1.15–1.44)	≤ 20%: 1.16 (1.04–1.28)	0.17
HF	> 20%: 1.22 (1.16–1.28)	≤ 20%: 1.20 (1.18–1.22)	0.54	> 20%: 1.21 (1.11–1.33)	≤ 20%: 1.20 (1.05–1.37)	0.92
DM	> 25%: 1.25 (1.23–1.28)	≤ 25%: 1.17 (1.16–1.19)	< 0.001	> 25%: 1.23 (1.13–1.35)	≤ 25%: 1.12 (0.97–1.30)	0.28
Trails	Clinical only: 1.28 (1.16–1.42)	Observations only: 1.20 (1.18–1.21)	0.21	Clinical only: 1.20 (1.07–1.36)	Observations only: 1.21 (1.10–1.34)	0.92

*P value for interaction; BB — beta-blocker; CI — confidence interval; DM — diabetes mellitus; HF — heart failure; IHD — ischemic heart disease
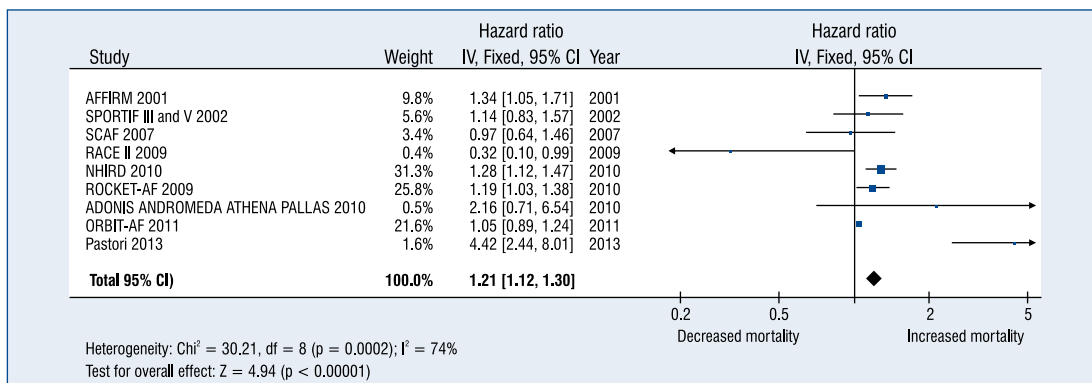
Supplementary Figure 2. Random effects model showing risk of cardiovascular disease with digoxin use (n = 9 studies).
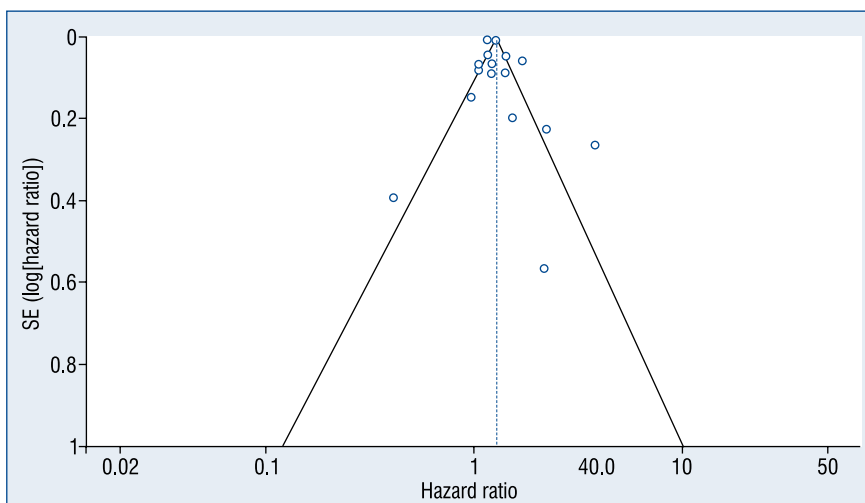
Supplementary Figure 3. Funnel plot showing publication bias.

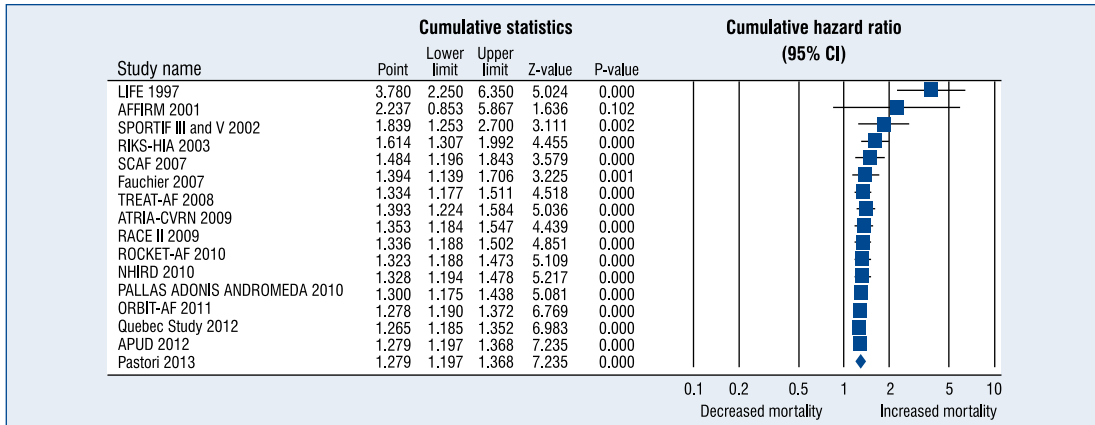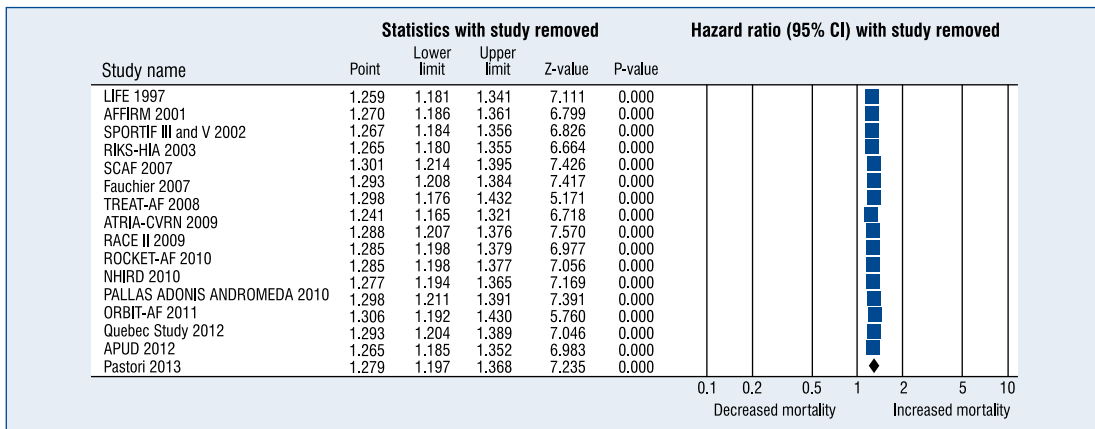| Study name	Cumulative statistics					Cumulative hazard ratio (95% CI)
	Point	Lower limit	Upper limit	Z-value	P-value	
LIFE 1997	3.780	2.250	6.350	5.024	0.000	
AFFIRM 2001	2.237	0.853	5.867	1.636	0.102	
SPORTIF III and V 2002	1.839	1.253	2.700	3.111	0.002	
RIKS-HIA 2003	1.614	1.307	1.992	4.455	0.000	
SCAF 2007	1.484	1.196	1.843	3.579	0.000	
Fauchier 2007	1.394	1.139	1.706	3.225	0.001	
TREAT-AF 2008	1.334	1.177	1.511	4.518	0.000	
ATRIA-CVRN 2009	1.393	1.224	1.584	5.036	0.000	
RACE II 2009	1.353	1.184	1.547	4.439	0.000	
ROCKET-AF 2010	1.336	1.188	1.502	4.851	0.000	
NHIRD 2010	1.323	1.188	1.473	5.109	0.000	
PALLAS ADONIS ANDROMEDA 2010	1.328	1.194	1.478	5.217	0.000	
ORBIT-AF 2011	1.300	1.175	1.438	5.081	0.000	
Quebec Study 2012	1.278	1.190	1.372	6.769	0.000	
APUD 2012	1.265	1.185	1.352	6.983	0.000	
Pastori 2013	1.279	1.197	1.368	7.235	0.000	
						0.1 0.2 0.5 1 2 5 10
						Decreased mortality Increased mortality

Supplementary Figure 4. Temporal effect shown by point estimates with addition of each study. Each point estimate is the cumulative hazard ratio in a way that the first study (LIFE 1997) showed only point estimate of LIFE 1997. The subsequent point estimate is the pooled estimate of LIFE 1997 and AFFIRM 2001, the third estimate is the pooled estimate of LIFE 1997, AFFIRM 2001 and SPORTIF III and IV 2002, and so on. Therefore, a decrement in the pooled estimate demonstrates a temporal decline in hazard ratio in published studies.

Study name	Statistics with study removed					Hazard ratio (95% CI) with study removed
	Point	Lower limit	Upper limit	Z-value	P-value	
LIFE 1997	1.259	1.181	1.341	7.111	0.000	
AFFIRM 2001	1.270	1.186	1.361	6.799	0.000	
SPORTIF III and V 2002	1.267	1.184	1.356	6.826	0.000	
RIKS-HIA 2003	1.265	1.180	1.355	6.664	0.000	
SCAF 2007	1.301	1.214	1.395	7.426	0.000	
Fauchier 2007	1.293	1.208	1.384	7.417	0.000	
TREAT-AF 2008	1.298	1.176	1.432	5.171	0.000	
ATRIA-CVRN 2009	1.241	1.165	1.321	6.718	0.000	
RACE II 2009	1.288	1.207	1.376	7.570	0.000	
ROCKET-AF 2010	1.285	1.198	1.379	6.977	0.000	
NHIRD 2010	1.285	1.198	1.377	7.056	0.000	
PALLAS ADONIS ANDROMEDA 2010	1.277	1.194	1.365	7.169	0.000	
ORBIT-AF 2011	1.298	1.211	1.391	7.391	0.000	
Quebec Study 2012	1.306	1.192	1.430	5.760	0.000	
APUD 2012	1.293	1.204	1.389	7.046	0.000	
Pastori 2013	1.265	1.185	1.352	6.983	0.000	
	1.279	1.197	1.368	7.235	0.000	
						0.1 0.2 0.5 1 2 5 10
						Decreased mortality Increased mortality

Supplementary Figure 5. One-study removal plot showing consistent estimates with removal of each study with each estimate. The corresponding hazard ratio in front of each study is the pooled hazard ratio after removal of that particular study. For example, the hazard ratio in front of LIFE 1997 is the pooled hazard ratio of all the other studies without including LIFE 1997. Such a plot demonstrates if one of the study has significant influential effect on the pooled estimates of meta-analysis. In a possible scenario, if a particular influential study is removed, the corresponding pooled hazard ratio may become insignificant. Here all of the corresponding pooled hazard ratios are significant, suggesting that none of the studies had particular influential effect on the pooled estimate.
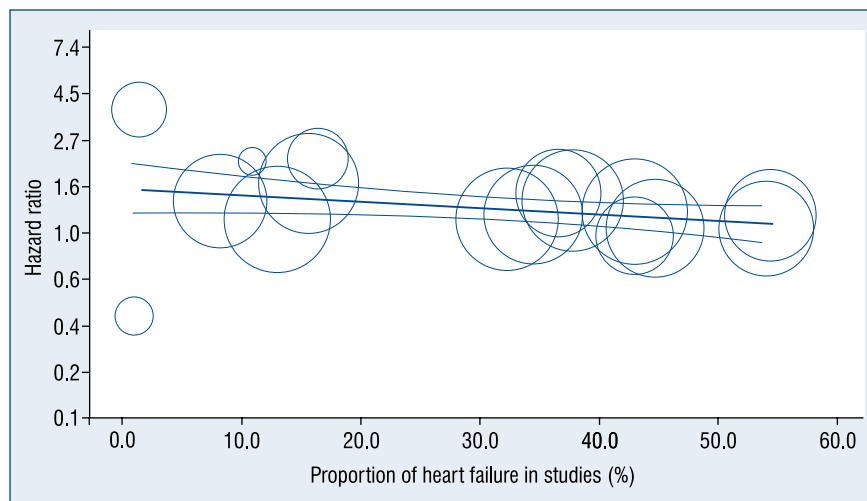
Supplementary Figure 6. Meta-regression plot showing a downsloping regression curve of hazard ratio across studies with increasing prevalence of heart failure patients for association of digoxin with mortality.